

Mathematical statistics recap for computing.

7 Limit Theorems

Motivation

For some new statistics, we may want to derive features of the distribution of the statistic.

When we can't do this analytically, we need to use statistical computing methods to *approximate* them.

We will return to some basic theory to motivate and evaluate the computational methods to follow.

7.1 Laws of Large Numbers

Limit theorems describe the behavior of sequences of random variables as the sample size increases ($n \rightarrow \infty$).

If X_1, \dots, X_n i.i.d.

Finite sample limit \rightarrow ① What is the distribution of $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

\rightarrow ② What is the approximate dsn as $n \rightarrow \infty$? Normal.

Often we describe these limits in terms of how close the sequence is to the truth.

How far is \bar{X} from μ ?

Statistic

True value we are estimating.

How to measure this distance? ex. $|\bar{X} - \mu|$ or $(\bar{X} - \mu)^2$

We can evaluate this distance in several ways.

Some modes of convergence – e.g.

– almost surely $P(\lim_{n \rightarrow \infty} X_n = X) = 1$.

– in probability $\forall \varepsilon > 0, \lim_{n \rightarrow \infty} P(|X_n - X| > \varepsilon) = 0$.

– in distribution $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$.

e.g. Laws of large numbers –

Weak LLN: Sample mean \bar{X}_n converges in probability to pop. mean μ .

Strong LLN: Sample mean \bar{X}_n converges a.s. to pop. mean μ .

What happens to sequences of r.v.s as n gets large (give us useful approximations!)

7.2 Central Limit Theorem

Theorem 7.1 (Central Limit Theorem (CLT)) Let X_1, \dots, X_n be a random sample from a distribution with mean μ and finite variance $\sigma^2 > 0$, then the limiting distribution of

$$Z_n = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \text{ is } N(0, 1). \quad (\text{converges in distribution})$$

i.e. $\bar{X}_n \xrightarrow{d} X, \quad X \sim N(\mu, \sigma^2/n).$

Interpretation:

The sampling distribution of the sample mean approaches a Normal distribution as the sample size increases.

Remember. Note that the CLT doesn't require the population distribution to be Normal.

8 Estimates and Estimators

Let X_1, \dots, X_n be a random sample from a population.

Let $T_n = T(X_1, \dots, X_n)$ be a function of the sample.

Then T_n is a "statistic"

and the pdf of T_n is called the "sampling distribution of T_n "

Statistics estimate parameters.

from sample

from population

Example 8.1

\bar{X}_n estimates μ

$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$ estimates σ^2

$S = \sqrt{S^2}$ estimate σ

Definition 8.1 An *estimator* is a rule for calculating an estimate of a given quantity.

Definition 8.2 An *estimate* is the result of applying an estimator to observed data samples in order to estimate a given quantity.

A statistic is a point estimator.

A CI is an interval estimator.

(if based on observed data, they are estimates)

We need to be careful not to confuse the above ideas:

\bar{X}_n function of r.v.'s \rightarrow estimator (statistic).

\bar{x}_n function of observed data (an actual #) \rightarrow estimate (sample statistic).

μ fixed but unknown quantities \rightarrow parameter

We can make any number of estimators to estimate a given quantity. How do we know the "best" one?

What are some properties we can use to say an estimator is

"better" than another one?

9 Evaluating Estimators

There are many ways we can describe how good or bad (evaluate) an estimator is.

9.1 Bias

Definition 9.1 Let X_1, \dots, X_n be a random sample from a population, θ a parameter of interest, and $\hat{\theta}_n = T(X_1, \dots, X_n)$ an estimator. Then the *bias* of $\hat{\theta}_n$ is defined as

$$\text{bias}(\hat{\theta}_n) = E[\hat{\theta}_n] - \theta.$$

Definition 9.2 An *unbiased estimator* is defined to be an estimator $\hat{\theta}_n = T(X_1, \dots, X_n)$ where

Example 9.1

Example 9.2

Example 9.3

9.2 Mean Squared Error (MSE)

Definition 9.3 The *mean squared error (MSE)* of an estimator $\hat{\theta}_n$ for parameter θ is defined as

$$\begin{aligned} MSE(\hat{\theta}_n) &= E \left[(\theta - \hat{\theta}_n)^2 \right] \\ &= Var(\hat{\theta}_n) + \left(bias(\hat{\theta}_n) \right)^2. \end{aligned}$$

Generally, we want estimators with

Sometimes an unbiased estimator $\hat{\theta}_n$ can have a larger variance than a biased estimator $\tilde{\theta}_n$.

Example 9.4 Let's compare two estimators of σ^2 .

$$s^2 = \frac{1}{n-1} \sum (X_i - \bar{X}_n)^2 \quad \hat{\sigma}^2 = \frac{1}{n} \sum (X_i - \bar{X}_n)^2$$

9.3 Standard Error

Definition 9.4 The *standard error* of an estimator $\hat{\theta}_n$ of θ is defined as

$$se(\hat{\theta}_n) = \sqrt{Var(\hat{\theta}_n)}.$$

We seek estimators with small $se(\hat{\theta}_n)$.

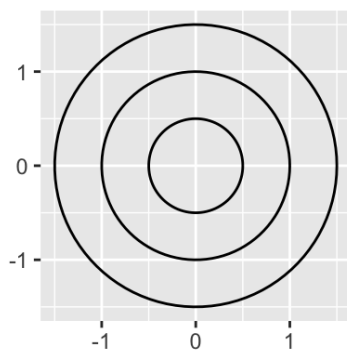
Example 9.5

10 Comparing Estimators

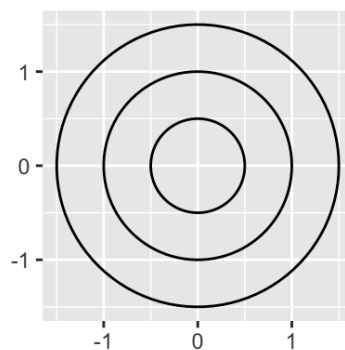
We typically compare statistical estimators based on the following basic properties:

- 1.
- 2.
- 3.
- 4.

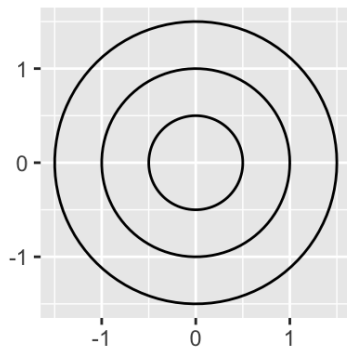
Unbiased and Inefficient



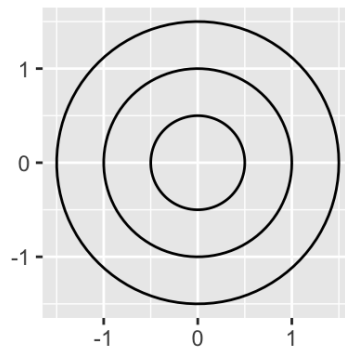
Biased and Inefficient



Biased and Efficient



Unbiased and Efficient



Example 10.1 Let us consider the efficiency of estimates of the center of a distribution. A **measure of central tendency** estimates the central or typical value for a probability distribution.

Mean and median are two measures of central tendency. They are both **unbiased**, which is more efficient?

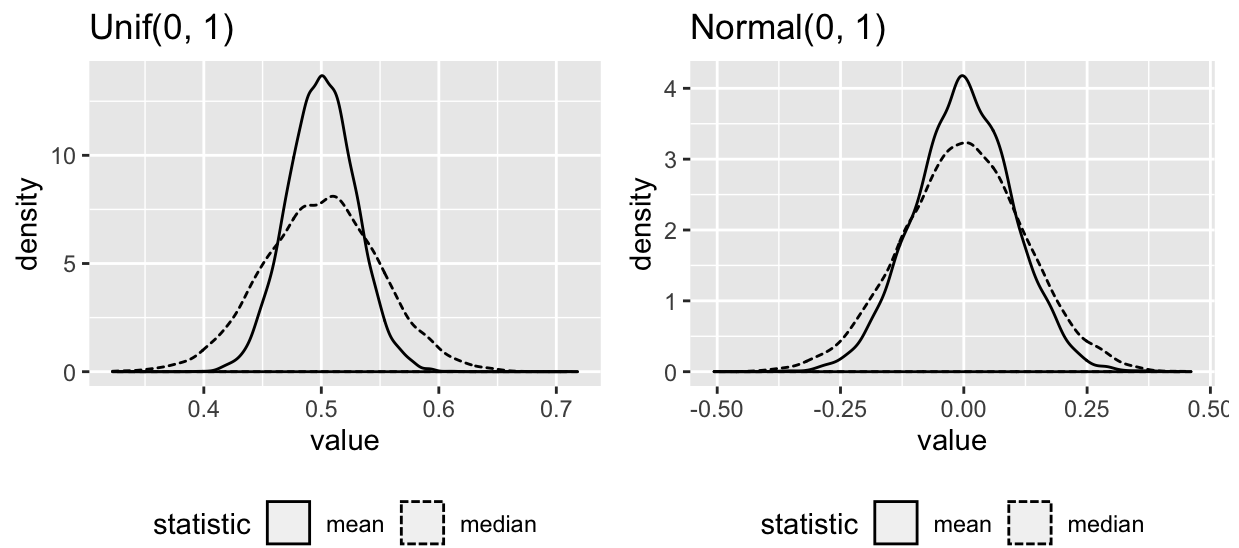
```
set.seed(400)

times <- 10000 # number of times to make a sample
n <- 100 # size of the sample
uniform_results <- data.frame(mean = numeric(times), median =
  numeric(times))
normal_results <- data.frame(mean = numeric(times), median =
  numeric(times))

for(i in 1:times) {
  x <- runif(n)
  y <- rnorm(n)
  uniform_results[i, "mean"] <- mean(x)
  uniform_results[i, "median"] <- median(x)
  normal_results[i, "mean"] <- mean(y)
  normal_results[i, "median"] <- median(y)
}

uniform_results %>%
  gather(statistic, value, everything()) %>%
  ggplot() +
  geom_density(aes(value, lty = statistic)) +
  ggtitle("Unif(0, 1)") +
  theme(legend.position = "bottom")

normal_results %>%
  gather(statistic, value, everything()) %>%
  ggplot() +
  geom_density(aes(value, lty = statistic)) +
  ggtitle("Normal(0, 1)") +
  theme(legend.position = "bottom")
```

Next Up In Ch. 5, we'll look at a method that produces *unbiased* estimators of $E(g(X))$!